Identification of Cancer: Mesothelioma's Disease Using Logistic Regression and Association Rule

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Abstract: Malignant Pleural Mesothelioma (MPM) or malignant mesothelioma (MM) is an atypical, aggressive tumor that matures into cancer in the pleura, a stratum of tissue bordering the lungs. Diagnosis of MPM is difficult and it accounts for about seventy-five percent of all mesothelioma diagnosed yearly in the United States of America. Being a fatal disease, early identification of MPM is crucial for patient survival. Our study implements logistic regression and develops association rules to identify early stage symptoms of MM. We retrieved medical reports generated by Dicle University and implemented logistic regression to measure the model accuracy. We conducted (a) logistic correlation, (b) Omnibus test and (c) Hosmer and Lemeshow test for model evaluation. Moreover, we also developed association rules by confidence, rule support, lift, condition support and deployability. Categorical logistic regression increases the training accuracy from 72.30% to 81.40% with a testing accuracy of 63.46%. The study also shows the top 5 symptoms that is mostly likely indicates the presence in MM. This study concludes that using predictive modeling can enhance primary presentation and diagnosis of MM.

Keywords: Logistic Regression, Mesothelioma, Predictive Modeling, Cancer Detection, Association Rules

Introduction

Malignant Pleural Mesothelioma (MPM) is a hostile tumor of mesothelial cells concomitant with preceding asbestos contact. With an amplified implementation of chemotherapy (Vogelzang et al., 2003; Zalcman et al., 2016) and a varied gamut of clinical examinations, precise prognostication is a crucial subject for individuals with MPM, doctors and scholars. However, MPM is an outstandingly different ailment. Staging system (Pass et al., 2016), challenging primary tumor identification process (Gill et al., 2016; Frauenfelder et al., 2011) and distinct biology (Bueno et al., 2016), impedes accurate prediction of MM. This fatal disease affects about two individuals per million per annum in a general population (McDonald and McDonald. 1996). Comparatively industrialized nations are affected more by MM (Spirtas et al., 1986; Peto et al., 1995; Leigh et al., 1991) due to higher exposure to asbestos (Metintas et al., 2008). The primitive symptoms of MM such as (a) puffing, (b) dyspnea, (c) respiratory complications, (d) pain in the chest or abdomen, (e) fever and night sweats, (f) pleural effusion, (g) fatigue and

(h) muscles weakness does not trigger a doctor to conduct a diagnosis of mesothelioma (MN, 2018) on time.

Predictive analytics can assist in the early discovery of diseases (Choudhury and Greene, 2018; Choudhury and Khan, 2018; Choudhury and Wesabi, 2018). However, distracted symptoms and various malignancies implicating the same tumor or cancer site may lead to a significant fraction of misclassifications leading to poor prediction accuracy. Coalescing evidence from various indicators using data mining techniques, such as Decision Tree (DT), Artificial Neural Network (ANN), Support Vector Machine (SVM) and other classifiers can enhance classification accuracy (Choudhury and Greene, 2018). Regrettably, each method inherits its limitations (Choudhury and Greene, 2018). For example, Random forest, decision tree and tree classifiers either tend to overfit (Choudhury and Khan, 2018) or fail to converge a large dataset (Choudhury and Wesabi, 2018).

In our analysis, we study the classification accuracy of the logistics regression model and compare its testing and training accuracy. Consecutively, our study proves that a logistic regression model gives better prediction than that of a base model.



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Methodology

Data Description

This study uses the patient's medical reports generated by Dicle University. The dataset contains 34 attributes, one binary response variable and 324 instances. Figure 1 shows the distribution of healthy and mesothelioma patients, where red bars represents mesothelioma patients and blue symbolizes healthy patients.

The dataset consists of 41% females and 59% males. Figure 2 shows the distribution of gender and health outcome.

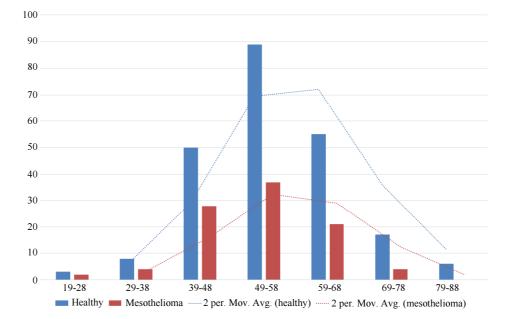


Fig. 1: Age versus health outcome

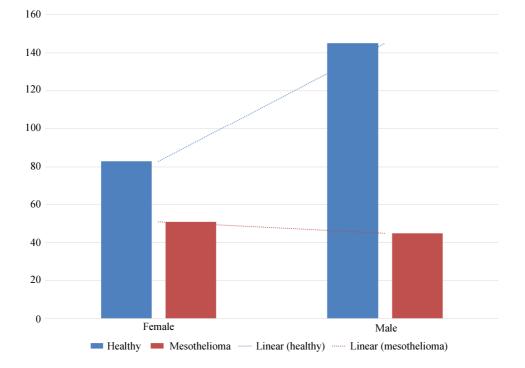
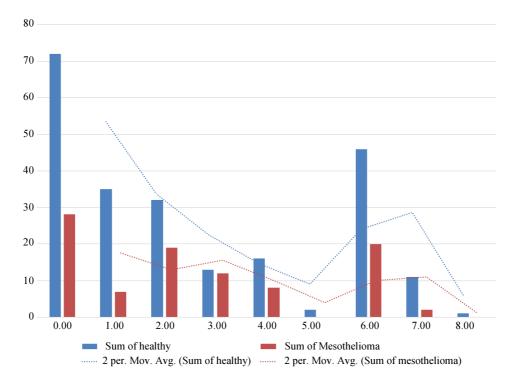


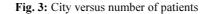
Fig. 2: Gender versus health outcome

The patients involved in this study belong to 9 different cities (0 through 8). Figure 3 shows the distribution of patients across cities. The x-axis represents the cities and Y-axis is the count of patients. City 0 consists of 22% healthy (h) and 9% mesothelioma (m) patients, city 1 consists of 11% h and 2% m, city 2 consists of 10% h and 6% m, city 3 consists of 4% h and

4% *m*, city 4 consists of 5% *h* and 2% *m*, city 5 consists of 1% *h* and no *m*, city 6 consists of 14% *h* and 6% *m*, city 7 consists of 3% *h* and 1% *m* and 1 individual from city 8 was found to be *h*.

Figure 4 shows the asbestos exposure for each city. It can be observed that the patients from city 0 have the highest asbestos exposure.





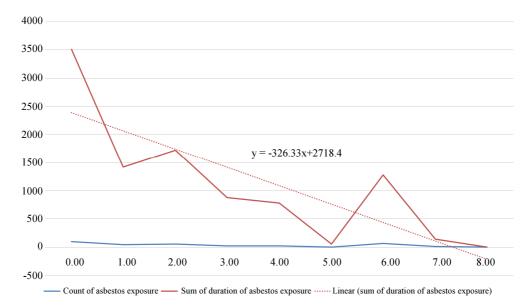


Fig. 4: City versus Asbestos exposure

Table 1: Data partitioning (training and testing dataset)

	Dataset partition	Dataset partitioning summary		
	N	Percent		
Training	220	67.90%		
Testing	104	32.09%		
Total	324	100.0%		

Data Preprocessing

We partitioned the data into training and testing dataset as shown in Table 1 and used training dataset for all the analysis. The training dataset was balanced using under-sampling (Parodi *et al.*, 2015) method.

Logistic Regression

Logistic regression is a widely used statistical technique (Tin, 1995). However, its typical use involves situations in which the outcome variable is continuous. Many situations in data analysis involve predicting the value of a nominal or an ordinal categorical outcome variable. In a situation in which we have a nominal categorical outcome variable, we use binary or multinomial logistic regression (Tin, 1995). In this study, we implement binomial logistic regression. Logistic regression is designed to use a mix of continuous and categorical predictor variables to predict a nominal categorical dependent variable. It does not directly predict the values of the dependent variable. Instead, the logistic equation predicts the odds of the event of interest occurring.

The general equation for logistic regression is:

$$Ln(Odds) = \alpha + B_1 X_1 + B_2 X_2 + \dots + B_k X_k$$
(1)

where, the terms on the right are the standard terms for the independent variable and the intercept in a regression equation (Tin, 1995), on the left side of the equation is the natural log of the odds (ln (Odds)) known as logit (Tin, 1995). The logit function is an Sshaped function. Logistic regression is managed by learning from the function as P(y = x). Y is a discrete value and X is a vector including discrete or continuous values. The algorithm estimates parameters from the training dataset. Logistic regression algorithm concludes probability and classifies the testing value by using threshold. Post optimizing the equation parameters, it can be employed to predict the output of testing dataset (Choudhury and Wesabi, 2018). Our study employs Categorical Regression which extends the regression model by quantifying categorical It (Tin, variables 1995). can also reduce multicollinearity among predictors, can model nonlinear relationships (Tin, 1995). Categorical regression maximizes the squared correlation between

the transformed dependent variable and the linear combination of the transformed predictors (Tin, 1995).

Categorical regression models make the same assumptions as linear regression models. Besides, categorical regression models assume that:

- There cannot be negative numbers in the data and all values must be integers (decimal digits are truncated)
- All nominal and ordinal variables should be coded so that their values are consecutive integers beginning with 1

In classification applications, calculating logistic dependencies between a single input and single target or class variable can be helpful. It determines the outright values of the logistic correlation concerning all predictors and all response variables. The logistic correlation is a statistical value between zero and one that conveys the métier of the logistic association between a single predictor and response variables. A value approaching one indicates strong relationship and value approaching 0 denotes weak or no relationship.

We calculate the absolute value of the logistic correlation between all predictors and response variable as presented in Table 2. Since "diagnosis method" can directly predict the presence or absence of MM, we exclude it from all further analysis.

Table 3 shows the dependent variable coding. It tells us that "healthy" is coded as 0 and 1 represents "mesothelioma."

We then measure the performance of the baseline model. Baseline model is a model that does not include the explanatory variables.

Explanatory variables, also known as independent or predictor variables, are factors that are operationalized and used in a regression to predict a given outcome (Leech *et al.*, 2015). The predictions of the baseline model are based purely on the frequency of occurrence of the category in the dataset.

Then we moved to the regression model that includes the explanatory variables and conducted The Omnibus tests. The Omnibus Tests of Model Coefficients measures the improvement of the new model over the baseline model (Leech *et al.*, 2015). It uses chi-square tests to see if there is a significant difference between the Log-likelihoods of the baseline model and the new model (Leech *et al.*, 2015). If the new model has a significantly reduced the Log-likelihood, then it suggests that the new model is elucidating more of the variance in the outcome and is an improvement. We calculated log likelihood and pseudo-R-square to determine the variation in the outcome. Consecutively, Hosmer and Lemeshow test (Leech *et al.*, 2015) was conducted to determine the goodness on the model.

Sl. No.	Input variable	Logistic correlation with "class of diagnosis."
1	Age	0.066
2	Gender	0.155
3	City	0.029
4	Asbestos exposure	0.079
5	Type of MM	0.134
6	Duration of asbestos exposure	0.069
7	Diagnosis method	1.000 (excluded from analysis)
8	Keep side	0.105
9	Cytology	0.029
10	Duration of symptoms	0.022
11	Dyspnea	0.026
12	Ache on chest	0.050
13	Weakness	0.060
14	Habit of cigarette	0.055
15	Performance status	0.039
16	White blood	0.050
17	Cell count (WBC)	0.052
18	Hemoglobin (HGB)	0.032
19	Platelet count (PLT)	0.065
20	Sedimentation	0.006
21	Blood Lactic Dehydrogenize (LDH)	0.014
22	Alkaline Phosphate (ALP)	0.041
23	Total protein	0.018
24	Albumin	0.041
25	Glucose	0.014
26	Pleural lactic dehydrogenize	0.036
27	Pleural protein	0.035
28	Pleural albumin	0.071
29	Pleural glucose	0.016
30	Dead or not	0.039
31	Pleural effusion	0.031
32	Pleural thickness on tomography	0.011
33	Pleural level of acidity (pH)	0.041
34	C reactive protein (CRP)	0.118

Table 2: Logistic correlation with the target variable

Table 3: Dependent variable encoding

Original Value	Internal value
Healthy	0
Mesothelioma	1

Performance Measures

When evaluating supervised training results, it is essential to check the performance of the training and testing through the accuracy and the AUC values. The performance measures come from equations based on the contingency matrix. This matrix is based on verifying combinations of true and false cases. The cases are true positive TP, true negative TN, false negative FN and false positive FP (Choudhury, 2018). Accuracy shows the percentage of correctly estimated true positive cases in the dataset. Overall prediction accuracy (Lotfi and Keshavarz, 2014) was used to identify the best fit model. Table 4: Association rule model setting

Maximum Number of Rules	10
Minimum Condition Support	0.05
Minimum Confidence	0.10
Minimum Rule Support	0.05
Minimum Lift	2.00
Maximum Number of Items in a Rule	10
Maximum Number of Items in a Condition	6
Maximum number of Items in a Prediction	3
Use only True Value for Flag Fields	True
Allow Rules without Conditions	False
Evaluation Measure Sorting the Rules	Confidence

Association Rule

An association rule is an implication expression of the form $X \rightarrow Y$, where X and Y are disjoint sets (McCormick and Salcedo, 2017). The strength of an association rule is measured concerning its support and confidence. Support (s) determines how often a rule applies to a given dataset, while confidence (c) determines how frequently items in Y appear in transactions that contains X (Leech *et al.*, 2015):

$$s(X \to Y) = (\sigma(X \cup Y)) / N \tag{2}$$

$$c(X \to Y) = (\sigma(X \cup Y)) / (\sigma(X))$$
(3)

Previously unknown associations in the medical domain have been identified with the help of association rule mining in the medical literature (Leech *et al.*, 2015). Association rule mining has been used to find disease-disease, disease-finding and disease-drug co-occurrences in electronic health record data (Hristovski *et al.*, 2001; Swanson, 1990). We investigate the factors which contribute to MM disease. In our study, association rule mining, a computational intelligence approach, is used to identify these factors.

Table 4 shows the association model used in this study.

Results

Logistic Regression

Table 5 shows the processing summary of the training dataset. It tells us that the analysis includes 220 instances and no patients have missing data.

Table 6 describes the baseline model. In this study, the baseline model always guesses '0' because most participants were not affected by MM. The overall percentage row shows that this approach to prediction is correct 72.3% 0 of the time.

Table 7 shows us the coefficient for the constant (B_0). According to this table, the base model with just the constant is a statistically significant predictor of the outcome (p<0.001).

Table 8 shows the omnibus tests outcome. It has three rows: (a) step, (b) block and (c) model. The Model row compares the new model to the baseline. The Step and Block rows are only essential if explanatory variables were added to the model in a stepwise or hierarchical manner. If we were building the model up in stages, then these rows would compare the Log-likelihoods of the newest model with the previous version to ascertain whether or not each new set of explanatory variables were triggering improvements. In our study, we have added all explanatory variables in one block and therefore have only one step.

The Sig. values are p<0.001, which indicates the accuracy of the model improves after adding the explanatory variables. Table 9 provides the log likelihood and pseudo-R2 values for the full model. The R2 values show the approximate variation in the outcome. We prefer to use the Nagelkerke's R2 which suggests that the model explains 43.50% of the variation in the outcome.

Table 5: Case processing summary Unweighted cases Percent Ν Selected cases Included in analysis 220 100.00% Missing cases 0 0 Total 220 100.00% Unselected cases 0.0 0 Total 220 100.00% Table 6. Classification table of the baseline model

				Predicted			
	Observed			Class_of_diagnosis			
				0.0	1.0	Perce	entage correct
Step 0	class_of_diagn	osis	0.0	159	0	100.0	0%
		~~		61	0	0	
	Overall Percentage					72.30%	
Tabla 7. V	ariables in the equation						
Table 7: V	ariables in the equation	on B	S.E.	Wald	df	Sig.	Exp(B)
Table 7: V Step 0	ariables in the equation		S.E. 0.151	Wald 40.463	df 1	Sig. 0.000	Exp(B) 0.384
Step 0		В -0.958			df 1		
Step 0	Constant	В -0.958		40.463	df 1 df		
Step 0 Table 8: O	Constant	B -0.958 coefficients	0.151	40.463	1		0.384
Step 0	Constant mnibus test of model	B -0.958 coefficients	0.151 Chi-squ	40.463	1 df		0.384 Sig.

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	del summary		10 1100				1 . D. G	
Step	-2 Log likelihood 180.838ª	-2 Log likelihood Cox and Snell R Square 180.838 ^a 0.301				Nagelkerke R Square		
$\frac{1}{a_{Eatimation t}}$	180.838" erminated at iteration number 8 be			logg the set 0.00	1	0.435		
Estimation to	erminatea at iteration number 8 be	cause log likelind	oa aecreasea t	by less than 0.00.	i percent			
Table 10: Ho	osmer and Lemeshow test							
Step	Chi-square			df			Sig.	
1	6.312			8			0.612	
Table 11: Cl	assification table of the new model	(training accurac		. 1				
			Predic	cted				
	class_of_diagnosis							
	Observed		0.0		0	Dore	antaga aarraat	
Step 1	Observed class of diagnosis	0.0	150	1.	9	94.3	centage correct	
Step 1	class_ol_diagnosis	1.0	32		9 19	94.5 47.5		
	Overall Percentage training		52	2	~	81.4		
	omparing training and testing classi	fication accuracy						
Partition	Training	_ · · ·	<i>co.</i> (esting			
Correct	179	81.3			56		63.46 %	
Wrong Total	41 220	18.6	4%0		38 04		36.54%	
Total	220			10	<i>J</i> 4			
Table 13: Ev	valuation matrix							
Partition	Training			T	esting			
Model	AUC	GIN	Ι		AUC		GINI	
Target	0.844	0.68	8	0.	0.613		0.277	
T.L. 14. D	1ab							
Table 14: Ru		Ma	ximum	Mean		Stand	ard deviation	
Measurement Condition Su		10.		9.20		0.75	ard deviation	
Confidence (10. 70.		66.67		3.23		
Rule Support			48	6.11		0.24		
Lift	2.11		38	2.25		0.11		
Deployability		3.	70	3.09		0.54		
a. Number of								
b. Number of	f Valid Events Data Source Records	s is 220						
Table 15: Top	five association rules sorted by confide	ence						
				Other Evaluation Statistics				
Condition		Pred.	Sorted by	Condition	Rule		Deployability	
	cness $7.600 \le \text{cell count (WBC)}$	Mesothelioma	confidence (% 70.37) Support (%) 8.33	Support (%) 5.86	Lift 2.38	<u>(%)</u> 2.47	
	$se \le 132.200 \ 1.760 \le pleural$	wesoulenoina	10.51	0.55	5.00	2.30	۲. ד.	
\leq cell count (W	sure dyspnoea weakness 7.600 VBC) < $11.2001.760 \le$ pleural	Mesothelioma	69.23	8.02	5.56	2.34	2.47	
	40 kness 7.600 ≤ cell count (WBC)) ≤ pleural albumin < 2.640	Mesothelioma	68.97	8.95	6.17	2.33	2.78	
dyspnoea weak	\leq pleural albumin < 2.640 dead or not	Mesothelioma	68.97	8.95	6.17	2.33	2.78	
dyspnoea weak	$cness 7.600 \le cell count (WBC)$ $0 \le pleural albumin < 2.640$	Mesothelioma	67.86	8.64	5.86	2.29	2.78	

As shown in Table 10, the Hosmer and Lemeshow test of the goodness of fit suggests the model is an

excellent fit to the data as p = 0.612 (>0.05). However, the chi-squared statistic on which it is based is dependent

on sample size, so the value cannot be interpreted in isolation from the size of the sample.

The Table 11 shows the training accuracy of the model that includes all the explanatory variables. The new model is now correctly classifying the outcome for 81.40% of the cases compared to 72.30% in the baseline or null model. The classification accuracy for class 1 increased from zero to 47.50%.

Table 12 compares the testing and training accuracy of the logistic regression model. Logistic regression gives a classification training accuracy of 81.36% and testing accuracy of 63.46%. The significant difference between the training and testing accuracy is due the small sample size and data bias.

Table 13 compares the AUC and GINI of training and testing model.

Association Rule

Table 14 shows the association rule statistics.

The following Table 15 shows the top 5 association rules and their evaluation statistics.

The association rule shows that, "dyspnea", "weakness", "WBC count" and "pleural albumin" indicates the presence of Mesothelioma.

Discussion

MPM is a belligerent cancer that is arduous to diagnose, specifically at early stages. In most of the developing nations, the reserves for histopathological diagnosis of suspected cases are limited. MPM is also a sporadic disease first diagnosed in 1962 (Musk *et al.*, 2011), with a very low likelihood of an individual suffering from this type of cancer.

A study exhibits the survival experience of Australian individuals with MM and found that survival has improved for each decade from the 1960s through 2000s (Musk et al., 2011). The progressive advances in survival time was observed was plausibly due to enhanced prognosis with time which resulted in earlier presentation, diagnosis and improved treatment (Musk et al., 2011). A review of a rift of cases from the first and last decades of 1960s and 2000s respectively, showed that, although the time between first presentation and diagnosis of MM did not alter (Musk et al., 2011); rather, time between reported onset of symptoms and diagnosis did reduce significantly (63 days in the 1970s to 31 days in the 2000s). This insinuates that earlier diagnosis is merely due to patients' awareness of their symptoms, leading to earlier presentation in primary care or speedier referral by general practitioners to specialists. It is also possible that the improvement in diagnosing and treating MM in the early 90s burgeoned due to advancement in medical science or bias in reporting clinical trials and outcome.

In recent years, subjects suffering from the sarcomatoid subtype of MM were excluded from

published clinical trials of active treatment because they did not respond as well as those with epithelioid or biphasic subtypes (Musk *et al.*, 2011). Barring of these and older patients from clinical trials tends to bias the overall survival of any study's participants giving a wrong impression that the prognosis of MM is improving.

Longer survival times for females have been reported previously (Kanazawa *et al.*, 2006; Marinaccio *et al.*, 2007; Mirabelli *et al.*, 2009; Neumann *et al.*, 2004); However, the biological cause responsible are unknown. Some studies have proposed that the difference may be due to misclassification as peritoneal MM of other abdominal neoplasms in females (Marinaccio *et al.*, 2007; Mirabelli *et al.*, 2009) Italian National Mesothelioma Register (Marinaccio *et al.*, 2007) reported augmented survival times in females with peritoneal, but not pleural.

Despite increasing resources and treatment expenses of MM there have been only modest improvements in survival over the past 40 yrs. Population-based study shows that median survival overall is still limited to less than a year from the time of diagnosis.

Therefore, primary diagnosis and prevention remains the most urgent priority for MM. Predictive analytics has the potential to advocate primary diagnosis of MM, increase the likelihood to patient survival.

Conclusion

In this study we provide a prediction model using logistic regression to diagnose the presence of MM based on early stage symptoms.

We can infer from the results that logistic regression can improve primary diagnosis of mesothelioma disease and is a better approach than using no predictive model. The underfitting (high training accuracy and low testing accuracy) behavior of the logistic model was also observed during this study. This study identifies, "dyspnea," "weakness," "WBC count," and "pleural albumin" as the essential attributes that indicates the presence of mesothelioma disease.

However, our study is of limited statistical power to estimate sound results due to small sample size. Another, more general, limitation is that we cannot estimate the effect of age and gender as risk factors of developing an MPM.

Availability of Data and Material

All data analyzed during this study are included in this published article and its supplementary information files.

Funding Information

This study was not funded by any internal or external source.

Ethics

No formal ethics approval was required in this particular case. The dataset used is publicly available for all researchers and no human participation was required for this study. The authors declare that they have no competing interests.

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